Correcting for Measurement Error in Reporting of Episodically Consumed Foods When Estimating Diet-Disease Relationships

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OUTLINE

- Introduction: impact of dietary measurement error
- Regression calibration correction
- Challenges to analysis of episodically consumed foods
- Regression calibration model
- Simulation study:
	- whole grains vs colorectal cancer in men
- fish vs colorectal cancer in men
- Example: red/processed meat vs lung cancer in NIH-AARP Diet & Health Study
- Discussion

Impact of Measurement Error

- Food Frequency Questionnaire (FFQ) is instrument of choice in most studies in nutritional epidemiology
- FFQ is known to contain substantial measurement error, random and systematic
- Typically measurement error causes two things: bias in the estimated exposure effect (often leading to flattened or attenuated true slope in disease model)
	- loss of statistical power to detect exposure effect

Impact of Measurement Error

 \bullet <u>Disease model</u>: for disease outcome *D*, vector $\boldsymbol{T} = (T_1, ..., T_K)^t$ of true usual intakes, and vector $\boldsymbol{Z} = (Z_1,...,Z_L)^t$ of covariates

$$
\mathbb{E}(D|T,\boldsymbol{Z})=m(\alpha_0+\boldsymbol{\alpha}_T^t\boldsymbol{T}+\boldsymbol{\alpha}_Z^t\boldsymbol{Z})
$$

where m^{-1} .) is link function (e.g., logit)

 \bullet <u>Main assumption</u>: errors in reported intakes Q are nondifferential with respect to outcome D , i.e.

$$
\mathcal{F}(D|\bm{T},\bm{Q},\bm{Z})=\mathcal{F}(D|\bm{T},\bm{Z})
$$

 $\overline{}$ = Example: conditional distribution of reported intakes given true intakes is the same among cases and controls

Regression Calibration

 \bullet Disease model

$$
\mathbb{E}(D| \boldsymbol{T}, \boldsymbol{Z}) = m(\alpha_0 + \boldsymbol{\alpha}_T^t \boldsymbol{T} + \boldsymbol{\alpha}_Z^t \boldsymbol{Z})
$$

• Regression calibration: to a very good approximation

$$
\mathbb{E}(D|\boldsymbol{Q},\boldsymbol{Z})=m(\alpha_0+\boldsymbol{\alpha}_T^t\mathbb{E}(\boldsymbol{T}|\boldsymbol{Q},\boldsymbol{Z})+\boldsymbol{\alpha}_Z^t\boldsymbol{Z})
$$

• Intuition: substitution for unknown vector T its best prediction given the reported intakes Q and covariates Z

Regression Calibration

- In absence of gold standard, regression calibration predictors $\mathbb{E}(T_k|\mathbf{Q}, \mathbf{Z}), k = 1, ..., K$ are estimated using short-term reference measurements
- For continuous intake, reference measurements are required to satisfy classical error model

$$
R_{ij} = T_i + \epsilon_{ij}
$$

where errors ϵ_{ij} are additive, independent of true intake, errors in FFQ, and each other

• Then regression calibration predictor can be estimated as

$$
\mathbb{E}(R_{ij}|\boldsymbol{Q}_i,\boldsymbol{Z}_i)=\mathbb{E}(T_i|\boldsymbol{Q}_i,\boldsymbol{Z}_i)
$$

Regression Calibration

- Ideal reference measure
	- short-term 'recovery' biomarker
- Reference measure in reality
	- more extensive short-term dietary-assessment method such as 24HR or diary
- 24HR is of special interest because it is used in 2 largest cohorts, AARP and EPIC
- Distributions of nutrient intakes are typically rather skewed: classical error model for reference measure may not hold
- Remedy: transformation to a scale where classical error model holds

Intake of Episodically Consumed Foods

• Problem: short-term reference measure (e.g., 24HR) has spike at zero and skewed distribution of positive intake

Statistical Model: true usual intake

• For person i , day j , and intake T_{ij} of interest, let

$$
p_i = \mathbb{P}(T_{ij} > 0|i)
$$

denote *probability* to consume on any given day \bullet Let

$$
A_i = \mathbb{E}(T_{ij}|i; T_{ij} > 0)
$$

denote usual consumption *amount*

• Then usual intake, defined as $T_i = \mathbb{E}(T_{ij}|i)$, is given by

$$
T_i = \mathbb{E}(T_{ij}|i; T_{ij} > 0) \times \mathbb{P}(T_{ij} > 0|i) = p_i A_i
$$

Statistical Model: assumptions for reference instrument

• Conditional on (transformed) $X_i = (Q_i^t, Z_i^t)^t$

$$
\mathbb{P}(R_{ij}>0|\boldsymbol{X}_i)=\mathbb{P}(T_{ij}>0|\boldsymbol{X}_i)
$$

• For a monotone transformation $g(.)$ reference amount on transformed scale has classical measurement error

$$
g(R_{ij}|R_{ij} > 0) = \mu_{R_i} + \epsilon_{ij}, \epsilon_{ij} \sim N(o, \sigma_{\epsilon}^2)
$$

• Reference amount is unbiased on transformed scale:

 $\mathbb{E}\{g(R_{ij})|i, R_{ij} > 0\} = g(A_i)$

Statistical Model: part I

ñ **Part I – Probability to consume**

Logistic regression (mixed model)

$$
\mathbb{P}(R_{ij}>0|\boldsymbol{X}_i)=\mathbb{P}(T_{ij}>0|\boldsymbol{X}_i)
$$

$$
= H(\beta_{01} + \boldsymbol{\beta}_{X1}^t \boldsymbol{X}_i + u_{1i})
$$

where

 $H(v) = (1 + e^{-v})^{-1}$ is logistic function $u_{1i} \sim N(0, \sigma_{u_1}^2)$ is person-specific random effect allowing person's value to differ from that defined by covariates

Statistical Model: part II

• Part II – Amount on consumption day

- Linear regression (mixed model) on transformed scale

$$
g(R_{ij}|R_{ij}>0; \mathbf{X}_i)=\mu_{R_i}+\epsilon_{ij}
$$

$$
=\beta_{02}+\boldsymbol{\beta}_{X2}^t\boldsymbol{X}_i+u_{2i}+\epsilon_{ij}
$$

where

 $q(v) = (v^{\theta} - 1)/\theta$ – Box-Cox transformation $u_{2i} \sim N(o, \sigma_{u_2}^2)$ – person-specific random effect $\epsilon_{ij} \sim N(o, \sigma_{\epsilon}^2)$ – within-person random error

Statistical Model

• Two-part model

 $\mathbb{P}(R_{ij}>0|\boldsymbol{X}_i)=H(\beta_{01}+\boldsymbol{\beta}_{X1}^t\boldsymbol{X}_i+u_{1i})$ $g(R_{ij}|R_{ij}>0; \mathbf{X}_i) = \beta_{02} + \boldsymbol{\beta}_{X2}^t \mathbf{X}_i + u_{2i} + \epsilon_{ij}$

• Link

$$
(u_{1i}, u_{2i})^t \sim N(\mathbf{0}, \mathbf{\Sigma}), \, \mathbf{\Sigma} = \begin{pmatrix} \sigma_{u_1}^2 & \rho_{u_1, u_2} \sigma_{u_1} \sigma_{u_2} \\ \sigma_{u_2}^2 \end{pmatrix}
$$

- person-specific random effects are correlated

- covariates can be (partially) shared

Regression Calibration Model

• True usual intake

$$
T_i = H(\beta_{01} + \boldsymbol{\beta}_{X1}^t \boldsymbol{X}_i + u_{1i}) \times g^{-1}(\beta_{02} + \boldsymbol{\beta}_{X2}^t \boldsymbol{X}_i + u_{2i})
$$

• Regression-calibration predictor for transformed $h(T_i)$

 $\mathbb{E}[h\{H(\beta_{01}+\boldsymbol{\beta}_{Y1}^t\boldsymbol{X}_i+u_{1i})q^{-1}(\beta_{02}+\boldsymbol{\beta}_{Y2}^t\boldsymbol{X}_i+u_{2i})\}|\boldsymbol{X}_i]$

• Linear regression calibration:

- Monte Carlo estimation of regression calibration predictors by generating $\hat{\mathbf{u}} = (\hat{u}_{1i}, \hat{u}_{2i})^t \sim N(\mathbf{0}, \hat{\boldsymbol{\Sigma}})$, using estimated parameters $(\hat{\beta}_{01}, \hat{\beta}_{02}, \hat{\boldsymbol{\beta}}_{X1}, \hat{\boldsymbol{\beta}}_{X2})$ to calculate $h(\hat{T}_i)$ and regressing $h(\widehat{T}_i)$ on X_i

EATS: Design

- Men and women 20-70 years
- Nationally representative sampling of $12,615$ telephone numbers
- Approximately 1600 recruited
- Four 24HRs, one in each season
- After one year: DHQ about past year
- 886 respondents completed four 24HRs and DHQ

Simulation Study

• Idea: simulate data that are similar to reported intake of *whole grains* and *fish* in EATS

- transform FFQ using best Box-Cox transformation to approximate normality

- fit two-part model relating 4 24HRs to transformed FFQ, Q^* , and estimate model parameters
- generate $u_i = (u_{1i}, u_{2i}) \sim N(0, \Sigma), i = 1, ..., 20,000$
- generate $Q_i^* \sim N(\mu_{Q^*}, \sigma_{Q^*}^2), i = 1, ..., 20,000$
- $-$ generate two 24HRs for 1,000 subjects in calibration study
- $R_{ij} = \begin{cases} 0 \text{ with } \text{pr} = 1 p_i, \ p_i = H(\beta_{01} + \beta_{Q1}Q_i^* + u_{1i}) \\ g^{-1}(\beta_{02} + \beta_{Q2}Q_i^* + u_{2i} + \epsilon_{ij}) \text{ with } \text{pr} = p_i \end{cases}$

Simulation Study

- Generate $T_i = H(\beta_{01} + \beta_{Q1}Q_i^* + u_{1i})g^{-1}(\beta_{02} + \beta_{Q2}Q_i^* + u_{2i})$ $i = 1, ..., 20,000$, and transform to $h(T_i)$ using best Box-Cox transformation to approximate normality
- Generate binary outcome variable for colorectal cancer in men

$$
\mathbb{P}(D_i = 1|T_i) = H\{\alpha_0 + \alpha_1 h(T_i)\}
$$

where α_1 represents $\log RR = 0.5$ for increasing exposure from 10 % to 90 % of the true exposure distribution and $\alpha_0 = -3.05$ which corresponds to probability of 3% for a 60 y old man to get disease in general population within 10 years

Simulation Study

- Comparison of 4 different methods:
	- true exposure on transformed scale
	- FFQ-reported exposure on transformed scale
	- "conventional" regression calibration approach by using mean of 2 24HRs on transformed scale as reference instrumentsuggested regression calibration
- Since different methods lead to fitting risk model on different scales, RR is always calculated for the given increase in intake from a to $b = a + \Delta$, where a is equal to 10th percentile and b is equal to 90th percentile of true exposure on original scale

Simulation Study: Results

• True log RR for increase in *whole grain* intake from 0.25 to 2.85 pyramid servings/day is equal to -0.74

Simulation Study: Results

• True log RR for increase in *fish* intake from 0.064 to 1.39 oz/day is equal to -0.69

NIH-AARP Diet & Health Study

- Prospective cohort of 567,169 men $&$ women aged 50-71 in 1995-96
- FFQ administered at baseline
- Calibration substudy of ~ 1000 men and ~ 1000 women with 2 24HRs and additional FFQ
- Analysis: association between red/processed meat and lung cancer for 349,148 men using Cox regression
- Confounders: age, BMI, smoking, physical activity, education, non-red/non-processed meat, fruit, total energy

NIH-AARP Diet & Health Study: Meat & Lung CA

Discussion

• New method addresses all of the challenges for modeling usual intake of foods and overcomes the limitations of conventional regression calibration

 Models intake as the product of probability to consume and consumption amount

 Allows for skewed distribution of reference consumption amount by transforming to a scale with classical error model

- Allows probability and amount to be correlated
- Uses rigorous regression calibration approach

Discussion

- Method is based on important assumptions that reference instrument correctly specifies probability of short-term fact of consumption and that, on appropriate scale, it follows classical measurement error for consumption amount
- Studies with unbiased biomarker (DLW) for energy expenditure have found bias in reporting of energy intake on 24HR - suggests systematic misreporting of at least some foods
- For foods reported with bias on 24HR, correction for measurement error using 24HR as reference instrument will be biased as well