

A Double Fixed Rank Kriging Approach to Spatial Regression Models with Covariate Measurement Error

Nickson Ning, Francis Hui, Alan Welsh

Research School of Finance, Actuarial Studies and Statistics



Australian
National
University

Ning, X., Hui, F.K.C., Welsh, A. (2023). A Double Fixed Rank Kriging Approach to Spatial Regression Models with Covariate Measurement Error. *Environmetrics*, **34**, e2771.
<https://doi.org/10.1002/env.2771>.

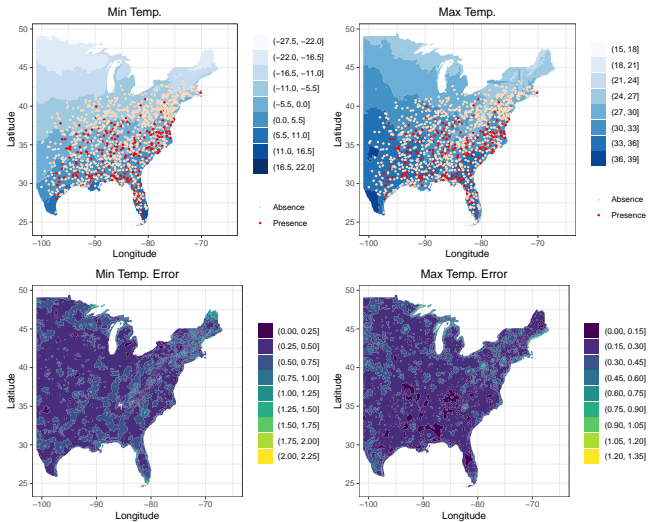


Carolina Wren



- ⊙ Bird
- ⊙ Cute
- ⊙ Provides motivating data for our methodology

Motivating data



- ⊙ Many disciplines e.g., environmental monitoring, ecology, epidemiology collect and analyse spatially indexed data
- ⊙ Often exhibit measurement error in the covariates

For $i = 1, \dots, n$, let $y(\mathbf{s}_i)$ denote a response measured at spatial location $\mathbf{s}_i \in \mathbb{D}$ in some spatial domain \mathbb{D} .

$$f(y(\mathbf{s}_i) | \vartheta(\mathbf{s}_i), \phi) = \exp \{ [y(\mathbf{s}_i)\vartheta(\mathbf{s}_i) - a\{\vartheta(\mathbf{s}_i)\}] / \phi + c(y(\mathbf{s}_i), \phi) \}$$

$$g\{\mu(\mathbf{s}_i)\} = \beta_0 + \mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + H_\rho(\mathbf{s}_i)$$

$H_\rho(\mathbf{s})$ is an unknown spatial field capturing the (residual) spatial correlation in the responses.

Commonly used for spatially indexed data (Diggle et al., 1998; Lindgren et al., 2011; Datta et al., 2016).

For observation $i = 1, \dots, n$ and covariate $j = 1, \dots, J$,

$$w_j(\mathbf{s}_i) = x_j(\mathbf{s}_i) + \epsilon_{ij}, \quad (1)$$

$$\epsilon_{ij} \stackrel{ind.}{\sim} N(0, \sigma_j^2).$$

Well-known that naively using $w_j(\mathbf{s}_i)$ instead of $x_j(\mathbf{s}_i)$ causes attenuation/bias towards zero (see e.g., Wald (1940)).

- ⊙ Many ways to model $H_\rho(s_i)$ (Albert and McShane, 1995; Cressie and Johannesson, 2008)
- ⊙ Classical measurement error model (1) for non-spatial data is also well-studied (Carroll et al., 2006; Fuller, 2009), e.g., SIMEX
- ⊙ Both: Alexeeff et al. (2016); Zhang et al. (2016)
- ⊙ No need to know/estimate σ_j^2 : Huque et al. (2016)

For covariate $j = 1, \dots, J$, assume that $x_j(\mathbf{s}_i)$ are realisations from an underlying spatial field $G_j(\mathbf{s})$. Cressie and Johannesson (2008) Fixed-rank Kriging:

$$\begin{aligned}x_j(\mathbf{s}_i) &= G_j(\mathbf{s}_i) = \mathbf{B}_j^\top(\mathbf{s}_i)\boldsymbol{\theta}_j; & \boldsymbol{\theta}_j &\sim N(\mathbf{0}, \boldsymbol{\Sigma}_j) \\H_\rho(\mathbf{s}_i) &= \mathbf{B}_\rho^\top(\mathbf{s}_i)\boldsymbol{\theta}_\rho; & \boldsymbol{\theta}_\rho &\sim N(\mathbf{0}, \boldsymbol{\Sigma}_\rho).\end{aligned}\tag{2}$$

- ⊙ **First step:** For each covariate $x_j(\mathbf{s}_i)$, fit a basis function model (GAM/LMM) to the error-contaminated covariates $w_j(\mathbf{s}_i)$.

$$w_j(\mathbf{s}_i) = \mathbf{B}_j^\top(\mathbf{s}_i)\boldsymbol{\theta}_j + \epsilon_{ij}.$$

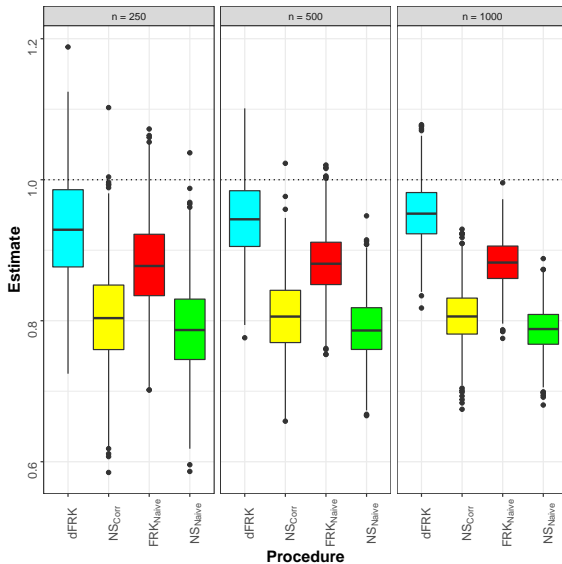
Using the model, get predictions $\hat{x}_j(\mathbf{s}_i) = \mathbf{B}_j^\top(\mathbf{s}_i)\hat{\boldsymbol{\theta}}_j$ for the underlying true covariates $x_j(\mathbf{s}_i)$.

- ⊙ **Second Step:** Fit the spatial GLMM (e.g., using a Laplace approximation), treating the $\hat{x}_j(\mathbf{s}_i)$'s as if they were the true underlying covariate values $x_j(\mathbf{s}_i)$.

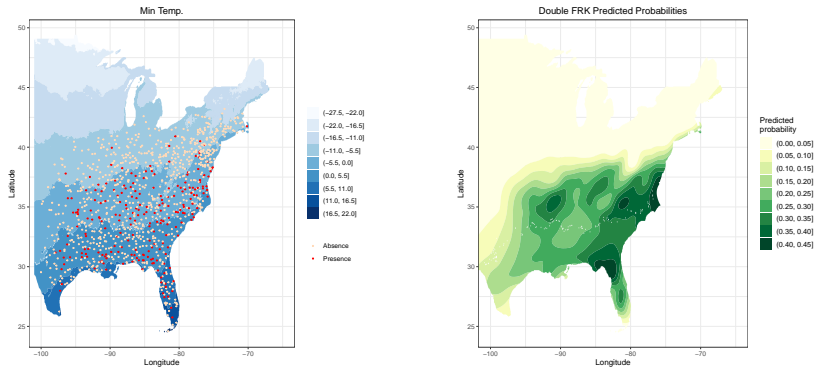
$$g\{\mu(\mathbf{s}_i)\} = \beta_0 + \hat{\mathbf{x}}(\mathbf{s}_i)^\top \boldsymbol{\beta} + \mathbf{B}_\rho^\top(\mathbf{s}_i)\boldsymbol{\theta}_\rho.$$

- ⊙ Simulated from spatial GLMM and covariate measurement error model
- ⊙ Compared slope estimates produced by different methods
- ⊙ *dFRK*: adjust for both covariate measurement error and residual spatial autocorrelation. *NS_{corr}*: No residual spatial autocorrelation adjustment. *FRK_{naive}*: No measurement error adjustment. *NS_{naive}*: no adjustment for either.

Simulations



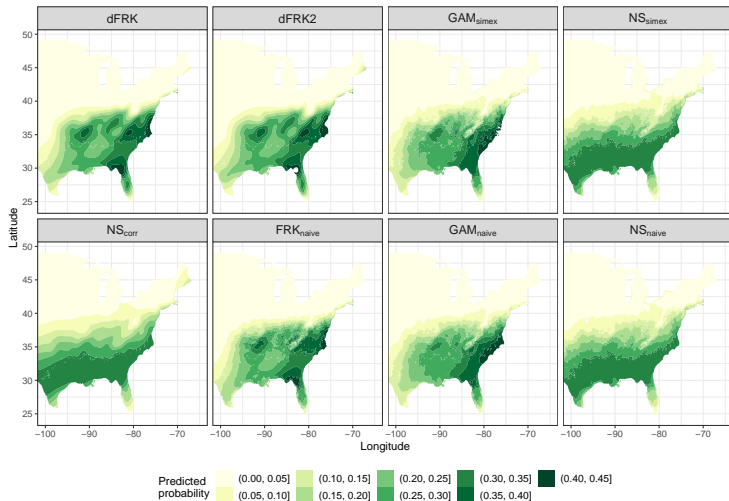
Motivating data revisited



Carolina Wren: Estimates

Spatial	M. Error	Estimates			Standard Error		
		Intercept	Temp.	Temp. ²	Intercept	Temp.	Temp. ²
FRK	FRK	-0.463	0.148	-0.0362	0.125	0.0247	0.0060
FRK ₂	FRK	-0.551	0.147	-0.0327	0.116	0.0244	0.0054
GAM	SIMEX	-0.607	0.192	-0.0224	0.320	0.1444	0.0065
None	SIMEX	-0.789	0.106	-0.0189	0.099	0.0190	0.0033
None	FRK	-0.793	0.106	-0.0186	0.098	0.0189	0.0033
FRK	None	-0.814	0.116	-0.0250	0.330	0.0869	0.0106
GAM	None	-0.769	0.119	-0.0212	0.258	0.1086	0.0058
None	None	-0.801	0.105	-0.0182	0.098	0.0188	0.0032

Carolina Wren: Out-of-sample Prediction



Carolina Wren: Out-of-sample Prediction

Spatial	M. Error	MSE	PD	AUC
FRK	FRK	0.1555	0.4764	0.6839
FRK ₂	FRK ₂	0.1562	0.4792	0.6830
GAM	SIMEX	0.1614	0.4882	0.6425
None	SIMEX	0.1581	0.4823	0.6444
None	FRK	0.1576	0.4808	0.6478
FRK	None	0.1570	0.4787	0.6673
GAM	None	0.1604	0.4853	0.6483
None	None	0.1580	0.4817	0.6468

Conclusion

- ⊙ Extend work of Huque et al. (2016) to a generalised linear mixed model framework
- ⊙ Deals with spatial correlation in responses and covariate measurement error simultaneously
- ⊙ No need for validation/replication data to estimate σ_j^2 : exploits the inherent spatial autocorrelation
- ⊙ Key assumption: covariates are realisations of some smooth underlying spatial field

THE
END

THANKS!