A Medley of Mixtures

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Summary

- 1 Motivating Example
- 2 Mixture models
- 3 Fish growth studies
- Mixtures of Mixed Models: Time Course Microarray
- 6 Clustering sugar cane SNP data
- 6 Acknowledgements
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Galaxy Data – Recession Velocities

Recession velocities (in 10^3 km/s) of 82 galaxies.



Galaxy Data - Fitted Normal Density

Single normal density based on sample mean and standard deviation



Is there any evidence of clustering?

Galaxy Data - Fitted Mixture Model

Mixture of normal densities with equal variances



Galaxy Data - Fitted Mixtures Models

Mixtures of normal densities



three components with unequal variances
 six components with equal variances

Mixture Models

- response y (uni/multivariate)
- explanatory variables **x**

K-component mixture

$$f(y \mid \boldsymbol{\Theta}, \mathbf{x}) = \sum_{k=1}^{K} \pi_k f_k(y \mid \boldsymbol{\theta}_k, \mathbf{x})$$

- *f_k* component densities (often same form)
- π_k component probabilities ($\sum_k \pi_k = 1$)
- θ_k component parameter vectors (some may be equal across components)

Galaxy Data - Fitted Mixtures Models

Mixtures of normal densities



----- : three components with equal variances ---- : six components with unequal variances

Estimation — EM algorithm

Likelihood — *n* observations

$$L(\boldsymbol{\Theta}) = \prod_{i=1}^{n} f(y_i \mid \boldsymbol{\Theta}, \mathbf{x}_i) = \prod_{i=1}^{n} \sum_{k=1}^{K} \pi_k f_k(y_i \mid \boldsymbol{\theta}_k, \mathbf{x}_i)$$

Estimation for finite mixture conveniently viewed as EM algorithm.

E-Step: Calculate component weights w_{ik} – the posterior probability that observation y_i comes from component k (useful for **clustering**):

$$w_{ik} = \frac{\pi_k f_{ik}}{\sum_{\ell} \pi_{\ell} f_{i\ell}}$$

M-step:

• component parameters: estimate $\widehat{\theta}_k$ from (y_i, \mathbf{x}_i) with weights w_{ik}

component proportions

$$\widehat{\pi}_k = \frac{\sum_{i=1}^n w_{ik}}{n}$$

Galaxy Data - Fitted Mixtures Models

Mixtures of normal densities



----- : six components with equal variances

Background

Fish ageing

Count rings on sectioned otolith (ear bone)



Courtesy of Irish Marine Institute

Background

Fish growth

Interested in the relationship between age and length or weight



Describing fish growth

Growth curves are typically 2-3 parameter non-linear models Most common is the von Bertalanffy:

$$\ell = \ell_{\infty} \left(1 - e^{-ka}
ight)$$

where

- *l*: length
- *a*: age
- ℓ_∞ : asymptotic length parameter
- k: growth rate parameter

Probabilistic (lognormal)

$$\ell_i = \ell_{\infty} \left(1 - e^{-ka_i} \right) e^{\varepsilon_i} \qquad \varepsilon_i \sim \mathsf{N}(0, \sigma_\ell^2)$$

Background

Fish growth

Fit a growth curve



Background

Fish growth

Sex-specific



Sex-specific growth

Standard practice is to discard juvenile data



Background

Sex-specific growth

Standard practice may not make most of the data:

- Focuses on known sexes only
- Uses a reduced sub-region of the age-length space
- May be uninformative on growth rate
- Likely very uninformative when third parameter introduced (non-zero y-axis intercept)

Sex-specific growth

Suggested alternative:

- Keep all data when fitting sex-specific growth curves
- Treat the sex of the juveniles as a classification problem
- Simultaneously estimate the juvenile sexes and growth curves

How?

Fish growth: mixture model

Outline:

$$f(\ell|a, \theta) = \pi_F f_F(\ell|a, \theta_F) + \pi_M f_M(\ell|a, \theta_M)$$

where

$$\pi_F = \Pr(S = F),$$

where S is the sex

$$f_{F}(\ell|\boldsymbol{a},\boldsymbol{\theta_{F}}) = \frac{1}{\ell\sigma_{F}\sqrt{2\pi}} \exp\left(-\frac{(\ln(\ell) - \ln(\nu(\boldsymbol{a},\boldsymbol{\theta_{F}})))^{2}}{2\sigma_{F}^{2}}\right)$$

Lognormal where v is the von Bertalanffy function

$$Z_i = \begin{cases} 1, & \text{if observation } i \text{ is female,} \\ 0, & \text{if observation } i \text{ is male.} \end{cases}$$

Note: Z is **partially classified** — we know the sex of **some** of the individuals

Example 1: separation

Data



Example 1: separation

Standard practice



Example 1: separation

Finite mixture model fit



Example 2: overlapping

Data



Example 2: overlapping

Status quo



Example 2: overlapping

Finite mixture model fit



Real Data



Traditional Modelling of Growth



Dashed: only known sex data; Solid: full knowledge fit

Mixture Modelling of Growth



Dashed: EM mixture model fit; Solid: full knowledge fit

Modelling of Maturity and Growth



Modelling of Maturity and Growth

Gender specific models for

- Growth length as a nonlinear model of age
- Maturity logit model for maturity (known gender) depending on age

Fitting strategies

- Separate fits for each model
- Joint fit missing gender estimation common to both models

Modelling of Maturity and Growth — Separate Fits



Dashed: EM fit; Solid: full knowledge fit

Modelling of Maturity and Growth — Joint Fit



Dashed: EM fit; Solid: full knowledge fit

Yeast data — time course microarray data





Data

Previous analyses

- Have traditionally been clustered using multivariate clustering methods, e.g. k-means clustering, hierarchical clustering, finite mixture models, etc.
- Problems with gene expression data?
 - High dimensionality;
 - Missing values;
 - Large amounts of measurement error;
 - Correlation between measurements made over time on same gene.
- Multivariate techniques have difficulties handling these issues.

Smoothing

- Assume that there exists some underlying function g(t) which generates the observed data.
- Observed data may contain a lot of measurement error/noise.

$$y_j = \underbrace{g(t_j)}_{\text{Signal}} + \underbrace{\varepsilon_j}_{\text{Noise}}$$

- Need to estimate smooth functions from noisy data.
- Use basis function expansions:

$$g(t) = \sum_{k=1}^{K} \beta_k \phi_k(t)$$

Basis functions

• Use *p*th degree truncated power basis (typically p = 1 or 2):

$$g(t_j) = \beta_0 + \beta_1 t_j + \ldots + \beta_p t_j^p + \sum_{\ell=1}^L \beta_{1\ell} (t_j - \kappa_\ell)_+^p,$$

$$\kappa_\ell = \ell \text{th knot and } (t_j - \kappa_\ell)_+ = \max(0, t_j - \kappa_\ell).$$



P-spline smoothing as a mixed model

- Represent P-spline smoothing as a linear mixed effects model.
- Mixed effects model has the form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \boldsymbol{\varepsilon}.$$

- For simplicity assume $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma_{\varepsilon}^2 \mathbf{I})$.
- For smoothing must also assume $\mathbf{u} \sim N(\mathbf{0}, \sigma_u^2 \mathbf{I})$.
- Estimates of β , σ_{ε}^2 , σ_{u}^2 and **u** determined using (RE)ML and BLUP.

Smoothing using mixed models: example

• Can smooth using a linear mixed effects model:

$$\mathsf{y} = \mathsf{X}eta + \mathsf{Z}\mathsf{u} + \epsilon$$

$$\boldsymbol{\beta} = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix} \text{ and } \mathbf{u} = \begin{pmatrix} \beta_{11} \\ \beta_{12} \\ \vdots \\ \beta_{1L} \end{pmatrix}$$

$$\mathbf{X} = \begin{pmatrix} 1 & t_1 \\ 1 & t_2 \\ \vdots & \vdots \\ 1 & t_n \end{pmatrix} \text{ and } \mathbf{Z} = \begin{pmatrix} (t_1 - \kappa_1)_+ & \cdots & (t_1 - \kappa_L)_+ \\ (t_2 - \kappa_1)_+ & \cdots & (t_2 - \kappa_L)_+ \\ \vdots & \ddots & \vdots \\ (t_n - \kappa_1)_+ & \cdots & (t_n - \kappa_L)_+ \end{pmatrix}$$

• Assume $\mathbf{u} \sim N(\mathbf{0}, \sigma_u^2 \mathbf{I})$ and $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma_{\varepsilon}^2 \mathbf{I})$.

Why bother?



- Very flexible.
- Computationally efficient.
- Can be fitted using readily available software, e.g. SAS, R, S-Plus, etc.

Gene expression clusters

- Want to cluster genes into groups exhibiting the same/similar expression profiles.
- Write the expression level for gene *i* in cluster *c* at time *j* as

$$y_{ij} = \mu_g(t_{ij}) + b_i + \varepsilon_{ij}, \quad j = 1, \ldots, n_i,$$

where $b_i \sim N(0, \sigma_{bc}^2)$ represent gene-specific shifts from the mean.

• Stack all data from genes in cluster c to get

$$\mathbf{Y}_{c} = \underbrace{\mathbf{X}_{c,s}\boldsymbol{\beta}_{c,s} + \mathbf{Z}_{c,s}\mathbf{u}_{c,s}}_{\mu_{c}(t)} + \mathbf{Z}_{c,b}\mathbf{b}_{c} + \boldsymbol{\varepsilon}_{c},$$

 $\mathbf{u}_{c,s} \sim N(\mathbf{0}, \sigma_{uc}^2 \mathbf{I}), \ \mathbf{b}_c \sim N(\mathbf{0}, \sigma_{bc}^2 \mathbf{I}), \ \boldsymbol{\varepsilon}_c \sim N(\mathbf{0}, \sigma_{\varepsilon c}^2 \mathbf{I}).$

Gene expression clusters

- In practice, do not know cluster membership.
- Assume **y**_i comes from a mixture of C clusters:

$$\mathbf{y}_i \sim \pi_1 \mathcal{N}(\mu_1(\mathbf{t}_i), \mathbf{V}_{i1}) + \pi_2 \mathcal{N}(\mu_2(\mathbf{t}_i), \mathbf{V}_{i2}) + \ldots + \pi_C \mathcal{N}(\mu_C(\mathbf{t}_i), \mathbf{V}_{iC})$$

where

$$\mu_{c}(\mathbf{t}_{i}) = \mathbf{X}_{i,s}\boldsymbol{\beta}_{c,s} + \mathbf{Z}_{i,s}\mathbf{u}_{c,s}$$

and $\mathbf{V}_{ic} = \sigma_{bc}^2 E_{n_i \times n_i} + \sigma_{\varepsilon c}^2 \mathbf{I}_{n_i \times n_i}$ • $\pi_1, \pi_2, \dots, \pi_C$ are mixing proportions such that $\sum_{c=1}^C \pi_c = 1$.

- Estimate π_1, \ldots, π_C , $(\boldsymbol{\mu}_1, \boldsymbol{V}_1), \ldots, (\boldsymbol{\mu}_C, \boldsymbol{V}_G)$.
- Obtain (posterior) probability that gene *i* is from cluster *c*.

Use EM algorithm.

EM algorithm



Cluster 3



EM algorithm



Cluster 3



Results

Results: BIC suggests 58 clusters; 6 example groups



Results



- GO terms:- Sterol transport and stress-response.
- Sterols important in many cellular processes (usually synthesised in the ER membrane).
- Anaerobic conditions: must be imported into the cell ⇒ sterol import genes activated.
- Other genes from Seripauperin family only activated under anaerobic conditions.

Sugar Cane



Motivation

- An allele is a particular form of a gene, e.g. the gene for eye colour has a number of alleles.
- Most organisms are diploid (2 sets of chromosomes).
- Sugarcane is polypoid (8 to 14 chromosomes) with individual alleles in varying numbers.
- Want to identify the many different alleles and associated genotypes/phenotypes.
- Can do this through the analysis of single nucleotide polymorphisms (SNPs).

SNPs

- SNPs occur during cell division, when cell divides in two by first copying its DNA.
- SNPs are mistakes that occur during the copying process i.e. changes that occur at a single base pair in DNA sequence.



- Frequency of a SNP base (A, T, C, G) at a locus determined by
 - the number of chromosomes carrying the gene;
 - the number of different alleles (or haplotypes);
 - the frequency of each allele possessing each SNP base.
- In sugarcane, the proportional frequencies of each SNP base varies depending on the number of alleles containing the SNP locus.
- Gives an indication of the number of allele haplotypes present for a gene.

Data — Spectra





 Homozygous individual with allele G (nucleotide)

 Heterozygous individual with some copies of allele C and some copies of allele G (G > C)

SNPs

Data — Idealised



Low Mass

Real Data

- How many clusters?
- What are the angles (dosages) and proportions?
- How to allocate the individuals?



Data — Illustrative



Want to develop a technique that can:

- determine the number of clusters present;
- determine the angles between the lines that represent each cluster to identify different genotypes;
- provide a probabilistic clustering to identify points that have high probability of belonging to a particular cluster (i.e. points that have a particular genotype) and those that are regarded as an unclear genotype.

Finite Mixture Models

- Have a *p*-length data vector $\mathbf{y}_i = (h.H, h.L)$ for each individual.
- Finite mixture models assume that the data come from a mixture of *G* clusters such that

$$f(\mathbf{y}_i; \boldsymbol{\theta}) = \sum_{g=1}^{G} \pi_g f_g(\mathbf{y}_i; \boldsymbol{\theta}_g), \qquad (1)$$

where $f(\mathbf{y}_i; \boldsymbol{\theta})$ is the density of the data, $f_g(\mathbf{y}_i; \boldsymbol{\theta}_g)$ is the gth component density and π_g are mixing proportions such that

$$\sum_{g} \pi_{g} = 1.$$

Usually assume

$$f_g(\mathbf{y}_i; \boldsymbol{\theta}_g) = N(\boldsymbol{\mu}_g, \boldsymbol{\Sigma}_g)$$
(2)

• Need to estimate $(\mu_1, \Sigma_1, \dots, \mu_G, \Sigma_G, \pi_1, \dots, \pi_{G-1})$ using EM algorithm.

Mclust



Linear Regression Lines

- Assume one of h.H/h.L is the response variable y_i and the other is the explanatory variable x_i.
- Fit a linear regression line through the origin

$$y_i = \beta_{1g} x_i + \varepsilon_i$$

to the data in each component.

Component densities now written as

$$f_g(y_i|\beta_{1g}x_i,\sigma_g^2) = N(\beta_{1g}x_i,\sigma_g^2),$$

where β_{1g} is the slope in the gth component and σ_g^2 is the variance.

• Need to determine estimates of β_{1g} , σ_g^2 (can be the same/different for each cluster) and $(\pi_1, \ldots, \pi_{G-1})$.

Results - Contig89b17



Model based clustering

Results - Contig2312b2



- Special case of Total Least Squares.
- Orthogonal regression line has the form

$$y_i = x_i \beta_{1g}$$

• Assumes both x and y are measured with error:

$$\begin{aligned} x_i &= x_i^* + \epsilon_i, \quad \mathsf{Var}(\epsilon) = \sigma_x^2 \\ y_i &= y_i^* + \tau_i, \quad \mathsf{Var}(\tau) = \sigma_y^2 \end{aligned}$$

- Orthogonal regression $\Rightarrow \sigma_x^2/\sigma_y^2 = \eta$ and independent.
- Suitable when both variables are linearly related and subject to error.
- Can fit a regression line to group parallel to the y-axis.



- Calculate $\hat{\beta}_{1g}$ using SVD.
- Need to find fitted values (\hat{x}_i, \hat{y}_i) .
- Find equation of line orthogonal to line with slope $\hat{\beta}_{1g}$ that goes through original data point (x_i, y_i) .
- Point at which this line and orthogonal regression line intersect gives fitted values:

$$\hat{x}_i = rac{y_i \hat{eta}_{1g} + x_i}{1 + \hat{eta}_{1g}^2}, \quad \hat{y}_i = \hat{x}_i \hat{eta}_{1g}$$

• Assume $\sigma_x^2 = \sigma_y^2 \Rightarrow$ overall estimate of σ_g^2 given by

$$\hat{\sigma}_g^2 = \frac{\sum_{i=1}^n (x_i - \hat{x}_i)^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2}{2(n-1)}$$

For clustering

$$f(x_i, y_i | \boldsymbol{\theta}_g) = \sum_{g=1}^{G} \pi_g f_g(x_i, y_i | \boldsymbol{\theta}_g)$$

• Component densities have form

$$f_g(x_i, y_i | \boldsymbol{\theta}_g) = N(\boldsymbol{\mu}_g, \boldsymbol{\Sigma}_g),$$

where

$$\boldsymbol{\mu}_{g} = \left(egin{array}{c} \hat{x}_{i} \ \hat{y}_{i} \end{array}
ight), \quad \boldsymbol{\Sigma}_{g} = \left(egin{array}{c} \hat{\sigma}_{g}^{2} & 0 \ 0 & \hat{\sigma}_{g}^{2} \end{array}
ight)$$

• For each component find \hat{x}_i , \hat{y}_i and $\hat{\sigma}_g^2$ as outlined in previous slide.

Results - Contig89b17



Results - Contig2312b2



Model choice & Use

• How many lines/groups?

Use BIC or similar approaches, as in standard model-based clustering.

- Constrained models lines at multiple of a common angle
- Which lines are present?

Estimated mixture proportions — give information on ploidy level and genotype distribution

Polar Coordinates: Results - Contig2312b2

$$r=\sqrt{x^2+y^2}$$
 $heta=rctan(y/x)$



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The End !!!

Thank you for your attention